

A Non-Convex Approach to Blind Calibration for Linear Random Sensing Models

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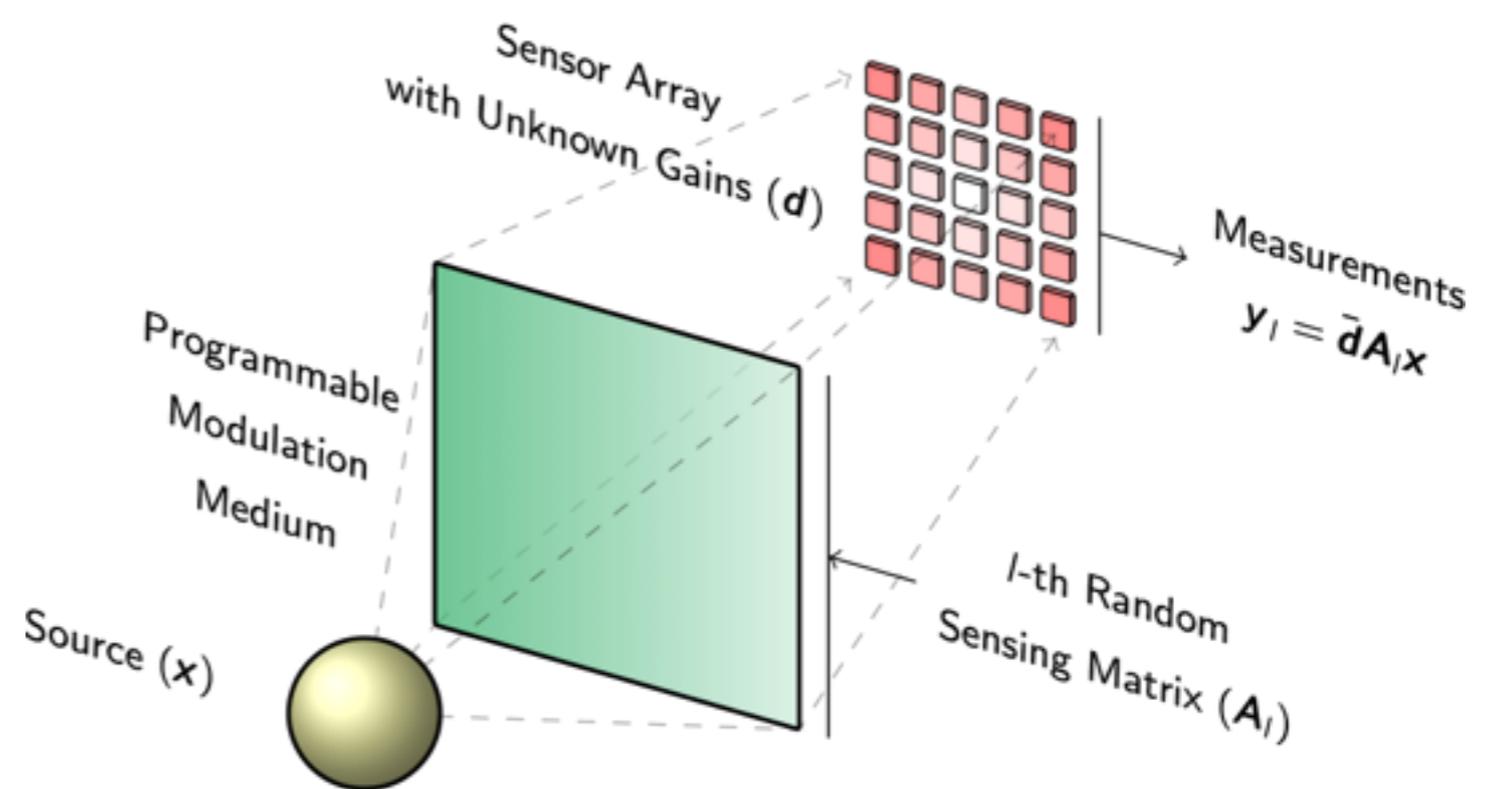
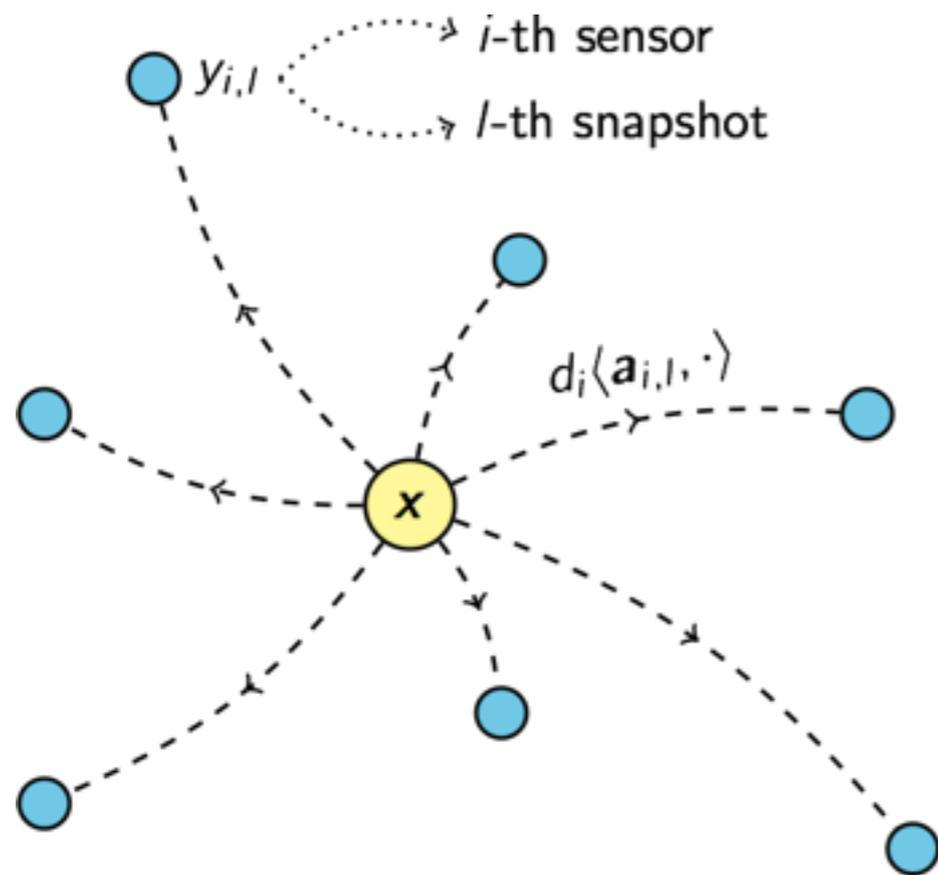
- **Introduction:** Blind Calibration and Random Linear Models
- **A Non-Convex Approach to Blind Calibration**
- **Solution by Projected Gradient Descent**
- **Global Convergence Guarantees** (even if non-convex!)
- **Experiments I:** Empirical Phase Transition
- **Experiments II:** Computational Imaging
- **Conclusion** and outlook

Random Linear Sensing Model

$$y_{i,l} = d_i \langle \mathbf{a}_{i,l}, \mathbf{x} \rangle, d_i > 0, \mathbf{x} \in \mathbb{R}^n$$

$$i = 1, \dots, m, l = 1, \dots, p \quad \equiv \quad \mathbf{y}^{(l)} = \text{diag}(\mathbf{d}) \mathbf{A}^{(l)} \mathbf{x}, \mathbf{d} \in \mathbb{R}_+^m, \mathbf{x} \in \mathbb{R}^n$$

$$l = 1, \dots, p, mp \geq n + m$$

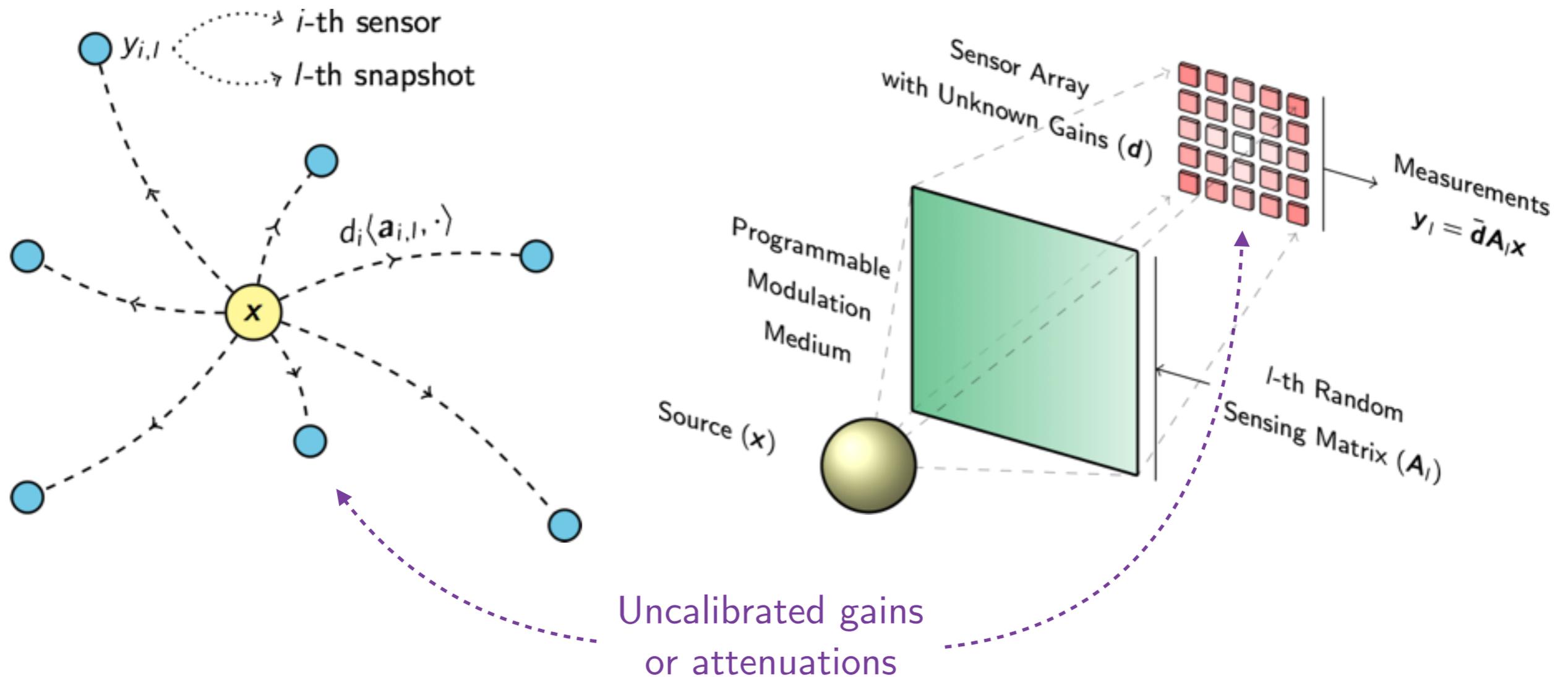


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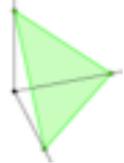
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- Which *algorithm* to *jointly* recover \mathbf{x} and \mathbf{d} ? Does it *provably* converge to the exact solution?
- How many *snapshots* and *measurements* suffice for an accurate recovery?
Sample complexity bound (mp) required with i.i.d. sub-Gaussian random vectors $\mathbf{a}_{i,l}$?

- We introduce the **Blind Calibration Problem**:

$$(\hat{\mathbf{x}}, \hat{\mathbf{d}}) = \underset{\xi \in \mathbb{R}^n, \gamma \in \Pi_+^m}{\operatorname{argmin}} \frac{1}{2mp} \sum_{l=1}^p \underbrace{\| \operatorname{diag}(\mathbf{d}) \mathbf{A}^{(l)} \mathbf{x} - \operatorname{diag}(\boldsymbol{\gamma}) \mathbf{A}^{(l)} \boldsymbol{\xi} \|_2^2}_{y^{(l)}}$$


(Scaled) probability simplex
Sum of Euclidean data fidelity terms

- The problem is clearly *non-convex* (indefinite Hessian matrix in general): the sensing model is *bilinear*, while the problem is *biconvex* (similar to *blind deconvolution*).
- The objective has minima in: $\{(\boldsymbol{\xi}, \boldsymbol{\gamma}) \in \mathbb{R}^n \times \mathbb{R}^m : \boldsymbol{\xi} = \frac{1}{\alpha} \mathbf{x}, \boldsymbol{\gamma} = \alpha \mathbf{d}, \alpha \in \mathbb{R} \setminus \{0\}\}$
- The constraint fixes one global minimiser: $(\mathbf{x}^*, \mathbf{d}^*) := \left(\frac{\|\mathbf{d}\|_1}{m} \mathbf{x}, \frac{m}{\|\mathbf{d}\|_1} \mathbf{d} \right)$
- Since the gains are *positive* and *bounded*, we let:

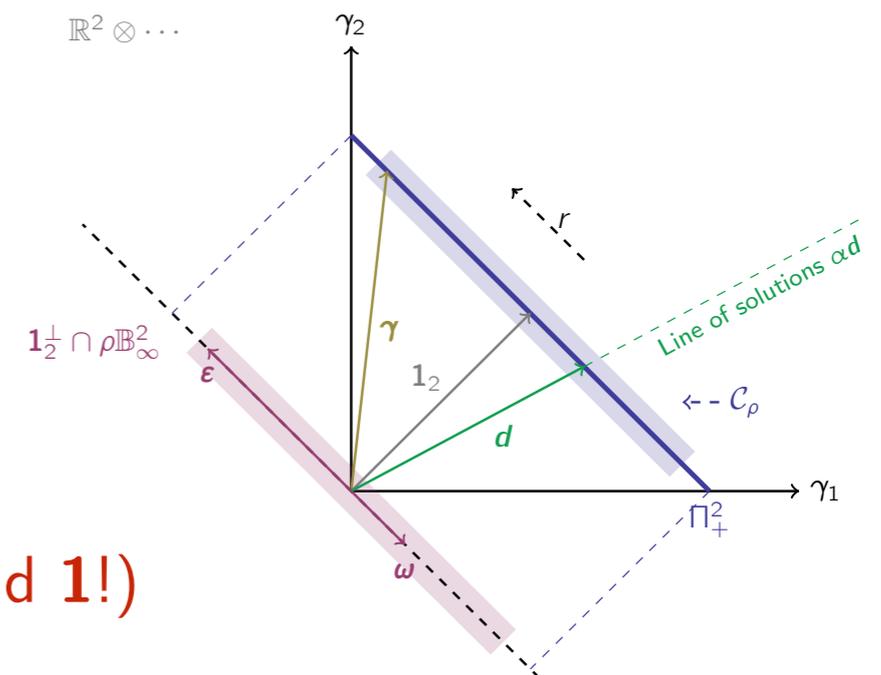
$$\boldsymbol{\gamma}, \mathbf{d}^* \in \mathcal{C}_\rho \subset \Pi_+^m, \mathcal{C}_\rho := \mathbf{1}_m + \mathbf{1}_m^\perp \cap \rho \mathbb{B}_\infty^m$$

$$\mathbf{d}^* = \mathbf{1}_m + \boldsymbol{\omega}, \boldsymbol{\omega} \in \mathbf{1}_m^\perp \cap \rho \mathbb{B}_\infty^m$$

$$\boldsymbol{\gamma} = \mathbf{1}_m + \boldsymbol{\varepsilon}, \boldsymbol{\varepsilon} \in \mathbf{1}_m^\perp \cap \rho \mathbb{B}_\infty^m$$

where:

$$\rho > \|\mathbf{d}^* - \mathbf{1}_m\|_\infty, \rho < 1 \quad (\text{Perturbation analysis around } \mathbf{1}!)$$



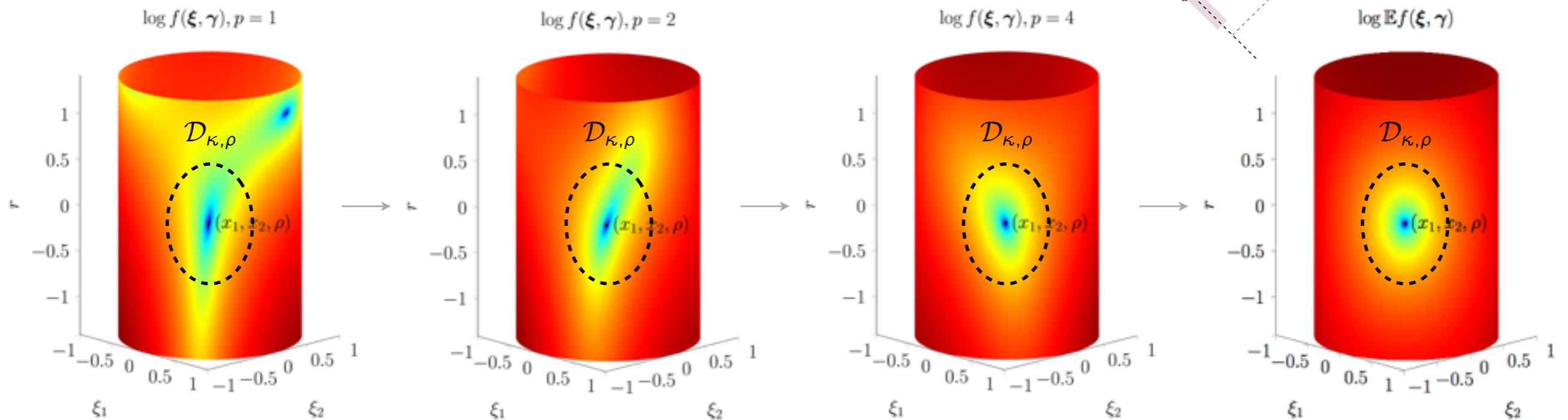
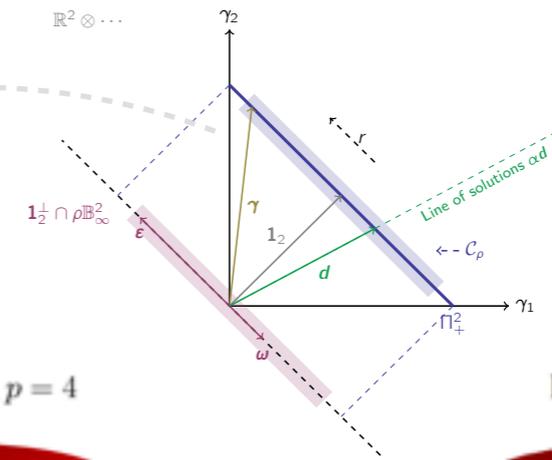
- The solution is obtained by *projected gradient descent*:

- 1: Initialise $\xi_0 := \frac{1}{mp} \sum_{l=1}^p (\mathbf{A}_l)^\top \mathbf{y}_l$, $\gamma_0 := \mathbf{1}_m$, $k := 0$.
- 2: **while** stop criteria not met **do**
- 3:
$$\begin{cases} \mu_\xi := \operatorname{argmin}_{v \in \mathbb{R}} f(\xi_k - v \nabla_\xi f(\xi_k, \gamma_k), \gamma_k) \\ \mu_\gamma := \operatorname{argmin}_{v \in \mathbb{R}} f(\xi_k, \gamma_k - v \nabla_\gamma^\perp f(\xi_k, \gamma_k)) \end{cases}$$
 {Line search in ξ, γ }
- 4: $\xi_{k+1} := \xi_k - \mu_\xi \nabla_\xi f(\xi_k, \gamma_k)$ {Signal Update}
- 5: $\gamma_{k+1} := \gamma_k - \mu_\gamma \nabla_\gamma^\perp f(\xi_k, \gamma_k)$ {Gain Update}
- 6: $\gamma_{k+1} := P_{C_\rho} \gamma_{k+1}$ {Projection on C_ρ }
- 7: $k := k + 1$
- 8: **end while**

$\nabla_\gamma^\perp f(\xi, \gamma) := P_{\mathbf{1}_m^\perp} \nabla_\gamma f(\xi, \gamma)$

- The chosen *initialisation* is *crucial*: in expectation (asymptotic p) it yields \mathbf{x} (unbiased estimator). For finite p it can be shown to lie *close* to the *global minimum*.
- Projection on C_ρ is only a technical requirement for proofs (not required in experiments).

- Consider a low-dimensional intuitive example for a random instance of the problem, at $n = 2, m = 2, \|\xi\|_2 = 1$, paramtrising $\gamma(r)$.



- To measure distances, we adopt the pre-metric:

$$\Delta(\xi, \gamma) := \|\xi - \mathbf{x}^*\|_2^2 + \frac{\|\mathbf{x}^*\|_2^2}{m} \|\gamma - \mathbf{d}^*\|_2^2.$$

- Thus, we define a *neighbourhood* of the global minimiser as:

$$\mathcal{D}_{\kappa, \rho} := \{(\xi, \gamma) \in \mathbb{R}^n \times \mathcal{C}_\rho : \Delta(\xi, \gamma) \leq \kappa^2 \|\mathbf{x}^*\|_2^2\}, \quad \rho \in [0, 1).$$

- Ideally: show via Hessian the local convexity of the problem in a given neighbourhood (for finite p , by *concentration of measure*).
- Simplification: *first-order properties* in the neighbourhood of the minimiser with i.i.d. sub-Gaussian random vectors. We need:
 1. **Initialisation**: fixes *radius* of neighbourhood, $(\xi_0, \gamma_0) \in \mathcal{D}_{\kappa, \rho}, \rho \in [0, 1)$
 2. **Regularity Condition**: developing the distance at iterate $k+1$,

$$\Delta(\xi_{k+1}, \gamma_{k+1}) = \Delta(\xi_k, \gamma_k) - 2 \underbrace{\left(\mu_\xi \langle \nabla_\xi f(\xi_k, \gamma_k), \xi_k - x^* \rangle + \mu_\gamma \frac{\|x^*\|_2^2}{m} \langle \nabla_\gamma^\perp f(\xi_k, \gamma_k), \gamma_k - g^* \rangle \right)}_{\text{Gradient Angle Part}} + \underbrace{\left(\mu_\xi^2 \|\nabla_\xi f(\xi_k, \gamma_k)\|_2^2 + \mu_\gamma^2 \frac{\|x^*\|_2^2}{m} \|\nabla_\gamma^\perp f(\xi_k, \gamma_k)\|_2^2 \right)}_{\text{Gradient Magnitude Part}} < \Delta(\xi_k, \gamma_k)$$

3. **Projection** on convex set \mathcal{C}_ρ ensures $\Delta(\xi_{k+1}, \gamma_{k+1}) \leq \Delta(\xi_{k+1}, \underline{\gamma}_{k+1})$

- Let's have a look at the gradient:

$$\nabla^\perp f(\xi, \gamma) = \frac{1}{mp} \sum_{l=1}^p \begin{bmatrix} \mathbf{A}_l^\top \text{diag}(\gamma) (\text{diag}(\gamma) \mathbf{A}_l \xi - \text{diag}(\mathbf{d}) \mathbf{A}_l \mathbf{x}) \\ \mathbf{P}_{1_m}^\perp \text{diag}(\mathbf{A}_l \xi) (\text{diag}(\gamma) \mathbf{A}_l \xi - \text{diag}(\mathbf{d}) \mathbf{A}_l \mathbf{x}) \end{bmatrix} \xrightarrow{p \rightarrow \infty} \frac{1}{m} \begin{bmatrix} \|\gamma\|_2^2 \xi - (\gamma^\top \mathbf{d}) \mathbf{x} \\ \|\xi\|_2^2 \boldsymbol{\varepsilon} - (\xi^\top \mathbf{x}) \boldsymbol{\omega} \end{bmatrix}$$

- To prove the minimum sample complexity that guarantees convergence, we need to verify two properties of the *initialisation* and the *neighbourhood* respectively.

Proposition (Initialisation Proximity). For any $\epsilon \in (0, 1)$ we have, with probability exceeding

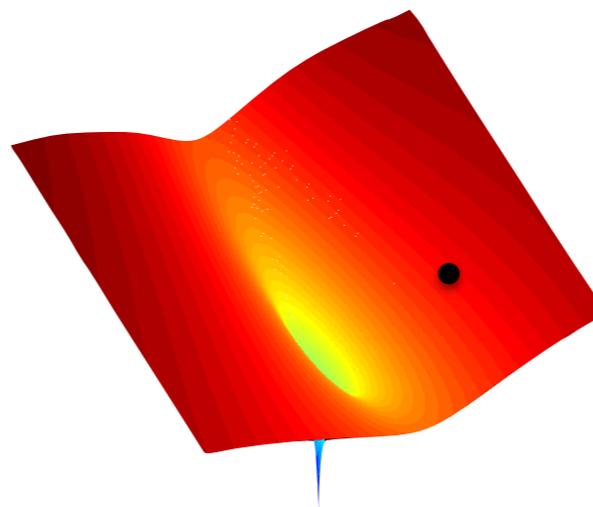
$$1 - Ce^{-c\epsilon^2 mp} - (mp)^{-t}$$

for some $C, c > 0$, that $\|\xi_0 - \mathbf{x}^*\|_2 \leq \epsilon \|\mathbf{x}^*\|_2$ provided $n \gtrsim t \log(mp)$ and

$$mp \gtrsim \epsilon^{-2}(n + m) \log\left(\frac{n}{\epsilon}\right).$$

Since $\boldsymbol{\gamma}_0 = \mathbf{1}_m$ we also have $\|\boldsymbol{\gamma}_0 - \mathbf{d}^*\|_\infty \leq \rho < 1$. Thus $(\xi_0, \boldsymbol{\gamma}_0) \in \mathcal{D}_{\kappa, \rho}$ with the same probability and $\kappa := \sqrt{\epsilon^2 + \rho^2} \leq \sqrt{2}$.

Initialise ...



Proposition (Regularity condition in $\mathcal{D}_{\kappa,\rho}$). For any $\delta \in (0, 1)$, $\rho \in [0, 1)$, $t > 0$, provided $\rho < \frac{1-2\delta}{9}$, $n \gtrsim t \log(mp)$, $p \gtrsim \delta^{-2} \log m$ and $\sqrt{mp} \gtrsim \delta^{-2}(n+m) \log(\frac{n}{\delta})$, with probability exceeding

$$1 - C [me^{-c\delta^2 p} + e^{-c\delta^2 \sqrt{mp}} + (mp)^{-t}]$$

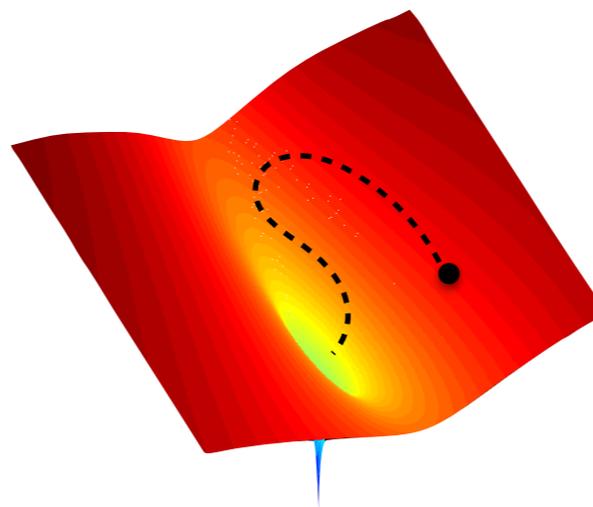
for some $C, c > 0$, we have that for all $(\xi, \gamma) \in \mathcal{D}_{\kappa,\rho}$,

$$\left\langle \nabla^\perp f(\xi, \gamma), \begin{bmatrix} \xi - x^* \\ \gamma - d^* \end{bmatrix} \right\rangle \geq \frac{1}{2} \eta \Delta(\xi, \gamma) \quad \text{(Bounded angle)}$$

$$\|\nabla^\perp f(\xi, \gamma)\|_2^2 \leq L^2 \Delta(\xi, \gamma) \quad \text{(Lipschitz gradient)}$$

for $\eta := 2(1 - 9\rho - 2\delta) > 0$, $L := 4\sqrt{2}[1 + \rho + (1 + \kappa)\|x^*\|_2]$.

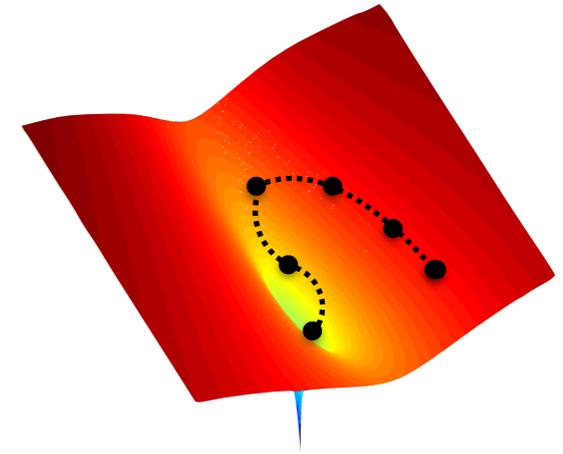
Initialise ...



... and Converge

$$\sqrt{mp} \gtrsim \delta^{-2}(n+m) \log\left(\frac{n}{\delta}\right)$$

- Under the previous conditions, we can bound the *error decay* of projected gradient descent when run in a neighbourhood of the global minimiser.
- The projection step serves to ensure *theoretically* that the neighbourhood does not change (*i.e.*, for the regularity condition)



Theorem (Provable Convergence to the Exact Solution). *Under the conditions of the previous Propositions we have that, with probability exceeding*

$$1 - C \left[m e^{-c\delta^2 p} + e^{-c\delta^2 \sqrt{m} p} + e^{-c\epsilon^2 m p} + (m p)^{-t} \right]$$

for some $C, c > 0$, our descent algorithm with $\mu_\xi := \mu, \mu_\gamma := \mu \frac{m}{\|\mathbf{x}^*\|_2^2}$ has error decay

$$\Delta(\boldsymbol{\xi}_k, \boldsymbol{\gamma}_k) \leq \left(1 - \eta\mu + \frac{L^2}{\tau}\mu^2\right)^k (\epsilon^2 + \rho^2) \|\mathbf{x}^*\|_2^2, (\boldsymbol{\xi}_k, \boldsymbol{\gamma}_k) \in \mathcal{D}_{\kappa, \rho}$$

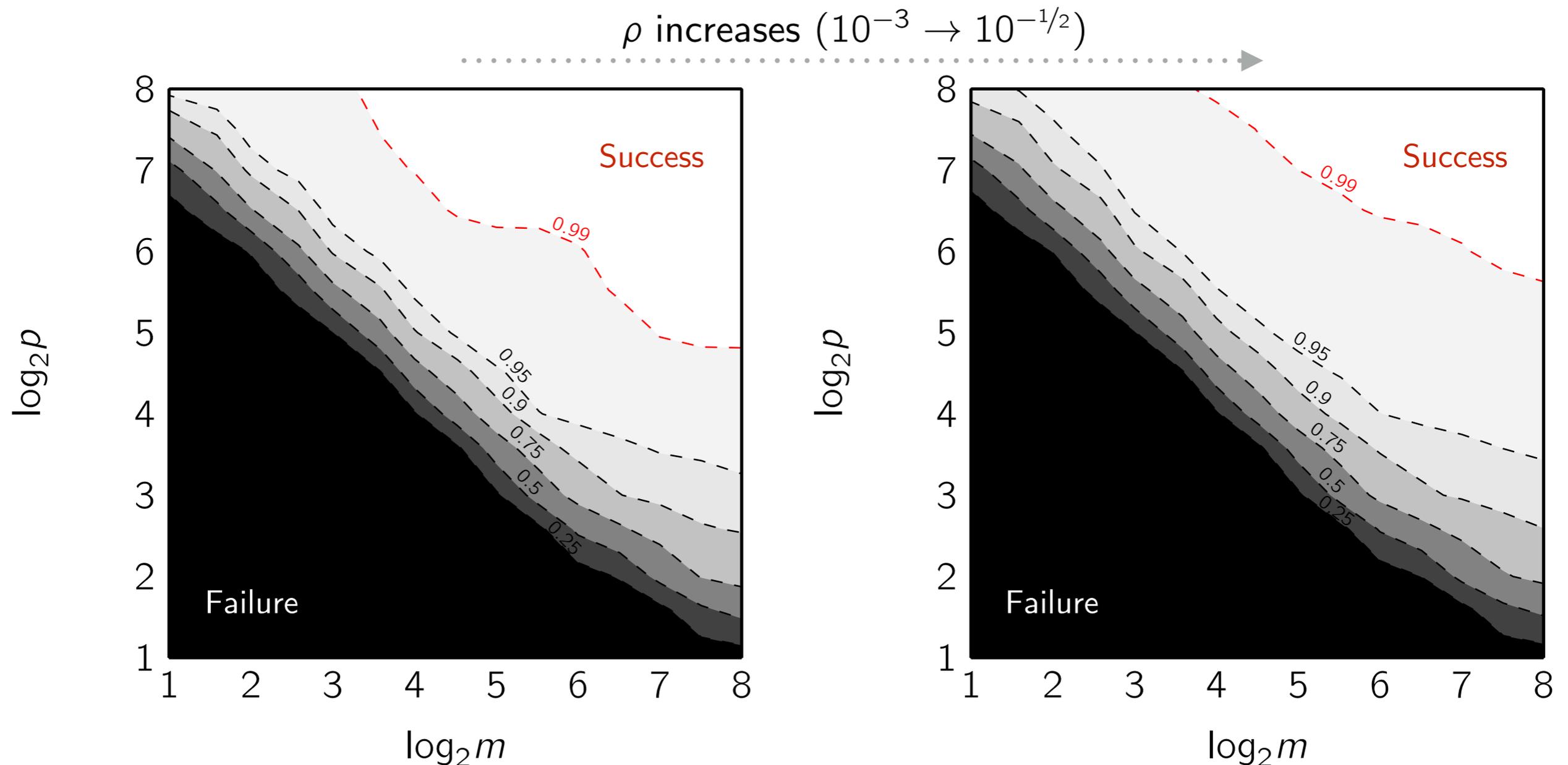
at any iteration $k > 0$ provided $\mu \in (0, \tau\eta/L^2)$, $\tau := \min\{1, \|\mathbf{x}^*\|_2^2/m\}$. Hence,

$$\Delta(\boldsymbol{\xi}_k, \boldsymbol{\gamma}_k) \xrightarrow[k \rightarrow \infty]{} 0.$$

- To test the problem's phase transition we measure the probability of successful recovery

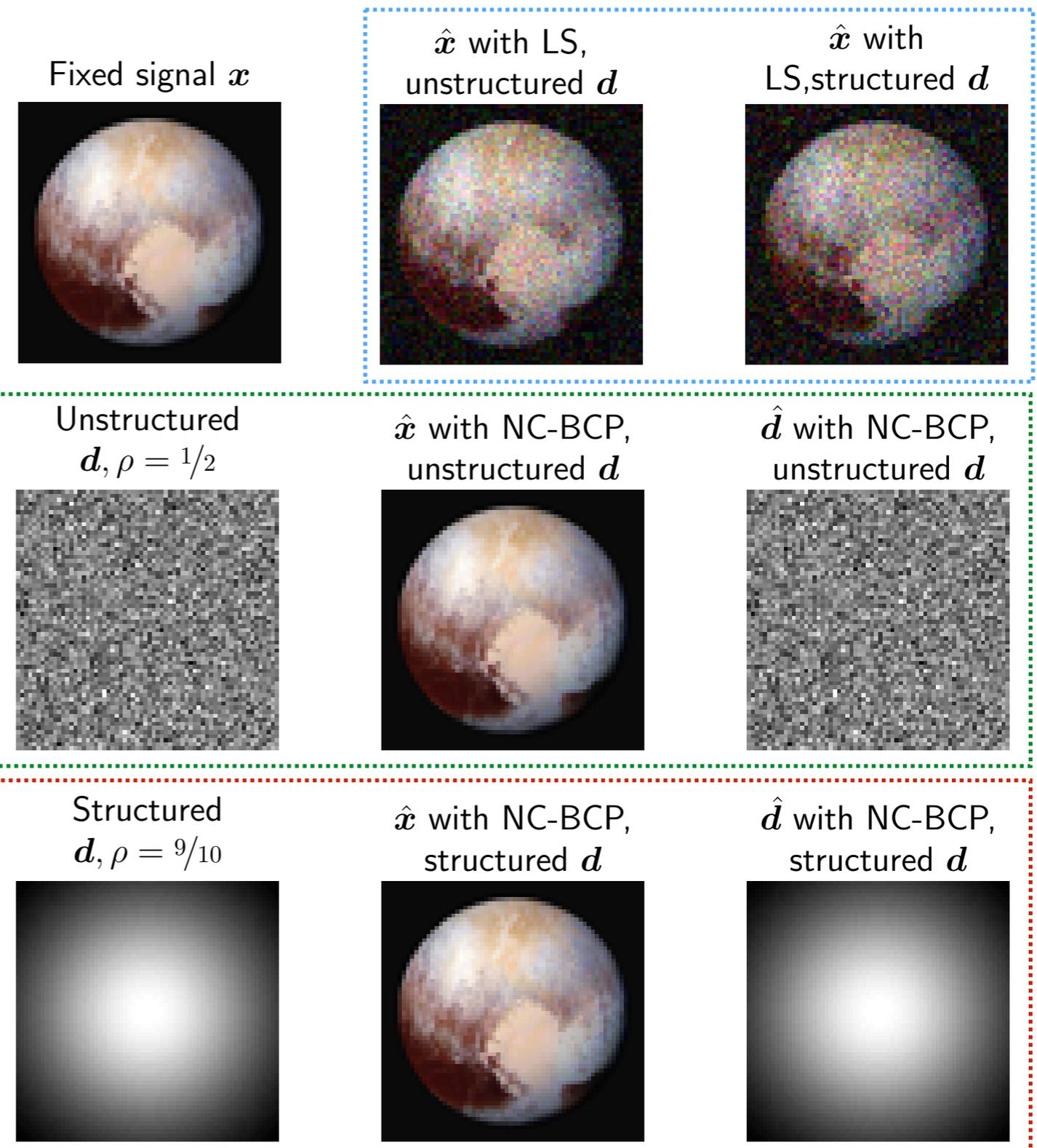
$$P_\zeta := \mathbb{P} \left[\max \left\{ \frac{\|\hat{\mathbf{d}} - \mathbf{d}^*\|_2}{\|\mathbf{d}^*\|_2}, \frac{\|\hat{\mathbf{x}} - \mathbf{x}^*\|_2}{\|\mathbf{x}^*\|_2} \right\} < \zeta \right], (\mathbf{x}^*, \mathbf{d}^*) \in R^n \times \mathcal{C}_\rho, n = 2^8$$

for 256 randomly generated problem instances (per point).



- An application to computational (compressive) imaging under calibration errors yields the following results for $p = 4$ snapshots when $m = n = 4096$.
- The achieved RMSE reads:

$$\max \left\{ \frac{\|\hat{\mathbf{d}} - \mathbf{d}^*\|_2}{\|\mathbf{d}^*\|_2}, \frac{\|\hat{\mathbf{x}} - \mathbf{x}^*\|_2}{\|\mathbf{x}^*\|_2} \right\} \approx -147.38 \text{ dB}$$
- The algorithm (NC-BCP) scales gracefully to very large values of n , contrarily to other approaches with guarantees.
- This experiment also converges with *fast* random matrices, such as a *subsampled random convolution* \mathbf{A}_l (not covered by current theory).



- We have shown that a *simple* application of gradient descent provably solves this bilinear inverse problem with sample complexity:

$$\sqrt{mp} \gtrsim (n + m) \log n, \quad p \gtrsim \log m, \quad n \gtrsim \log mp$$

- **Proved extensions** of this approach:
 - Blind calibration with *known* subspace signal/gain models (lower sample complexity).
 - Stability analysis w.r.t. additive noise.
 - Better sample complexity is possible (*linear* in number of unknowns).
- **Future developments:**
 - Extension to signal-domain *sparsity* via hard thresholding: reduces sample complexity (*i.e.*, blind calibration for compressed sensing); empirically shown, not yet proved.
 - Extension to related problems: blind calibration with complex gains and sensing matrices; blind deconvolution.
- Finding applications in which blind calibration of a sensor is critical and random measurements can be physically implemented in the sensing device.

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Thank you for your attention.

For any question or suggestion, contact us at:

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